2485

APPROXIMATION BY EXPONENTIALS. II.* EXAMINATION OF THE LIMITING CASES OF THE FUNCTION

 $u = A_1 \exp(\omega_1 x) + A_2 \exp(\omega_2 x)$

E.Slavíček

Department of Technical Physics, Institute of Chemical Technology, Prague 6

Received December 18th, 1970

The limiting cases of the function $u = A_1 \exp(\omega_1 x) + A_2 \exp(\omega_2 x)$, or $u = P_1 p_1^x + P_2 p_2^x$, lead for $\omega_2 \rightarrow \omega_1$, or $p_2 \rightarrow 0$, to relations of the type $u = (A + B_x) p_1^x$, or $u = A p_1^x + B$. The method of approximation by this empirical function is extended to the case of the data with unevenly spaced values of the argument.

The formulas for the sum of squares of deviations based on a sum of two exponentials may be divided into three groups:

Case I. Approximation formula of the general type

$$u_{i} = M_{1}m_{1}^{i} + M_{2}m_{2}^{i}, \qquad (1)$$

obtained from a more general formula

$$u = P_1 p_1^x + P_2 p_2^x \tag{1a}$$

assuming the uniform spacing of the experimental data. If $x = x_0 + ih$, where h is the increment of the argument, then

$$P_{j}p_{j}^{x} = P_{j}p_{j}^{(x_{0}+ih)} = P_{j}p_{j}^{x_{0}}(p_{j}^{h})^{i} = M_{j}m_{j}^{i}$$

for j = 1, 2; i = 0, 1, ..., R - 1. R is the number of measured pairs of data. Corresponding formula for the sum of squares, S, is

$$S = [y_i y_i] - M_1 [y_i m_1^i] - M_2 [y_i m^i].$$
⁽²⁾

Part I: This Journal 35, 2885 (1970).

Collection Czechoslov, Chem. Commun. /Vol. 37/ (1972)

The values of M_1 , M_2 are found from equations

$$M_{1}[m_{1}^{i}m_{1}^{i}] + M_{2}[m_{1}^{i}m_{2}^{i}] = [y_{i}m_{1}^{i}], \qquad (3)$$
$$M_{1}[m_{1}^{i}m_{2}^{i}] + M_{2}[m_{2}^{i}m_{2}^{i}] = [y_{i}m_{2}^{i}].$$

Case II. Approximation formula

$$u_{i} = M_{1}(m_{1}^{i} - m_{2}^{i}), \qquad (4)$$

which provides that the curve passes through the point $x_0 = 0$, $y_0 = 0$. The value of M_1 is found from equation

$$M_{1} = \frac{\left[y_{1}m_{1}^{i}\right] - \left[y_{1}m_{2}^{i}\right]}{\left[m_{1}^{i}m_{1}^{i}\right] - 2\left[m_{1}^{i}m_{2}^{i}\right] + \left[m_{2}^{i}m_{2}^{i}\right]^{2}},$$
(5)

The sum of squares is

$$S = [y_i y_i] - M_i [(m_1^i - m_2^i) y_i].$$
 (6)

Case III. Approximation formula

$$u_{i} = M_{1}m_{1}^{i} + (1 - M_{1})m_{2}^{i}, \qquad (7)$$

which provides that the curve passes through the point $x_0 = 0$, $y_c = 1$. The value of M_1 is found from equation

$$M_{1} = \frac{\left[(m_{1}^{i} - m_{2}^{i}) y_{i} \right] - \left[(m_{1}^{i} - m_{2}^{i}) m_{2}^{i} \right]}{\left[m_{1}^{i} m_{1}^{i} \right] - 2 \left[m_{1}^{i} m_{2}^{i} \right] + \left[m_{2}^{i} m_{2}^{i} \right]},$$
(8)

the sum of squares

$$S = [y_i y_i] - 2[m_2^i y_i] + [m_2^i m_2^i] - M_i[(m_1^i - m_2^i)(y_i - m_2^i)].$$
(9)

Clearly, for $m_2 = m_1 \text{ Eqs } (3)$ are identical and not suitable for evaluation of M_1 and M_2 . Eqs (5) and (8) as well give indefinite expressions for $m_2 = m_1$. It is therefore necessary to examine more closely the case $m_2 \rightarrow m_1$. One possible way is to substitute the relation $m_2 = m_1 (1 + x)$, express M_1 and S as functions of m_1 and x and to examine the limits for $x \rightarrow 0$. In case I, Eq. (3) has the solutions

$$M_1 = \Delta_1 / \Delta , \quad M_2 = \Delta_2 / \Delta , \qquad (10)$$

where

$$\Delta = \left[m_1^i m_1^i \right] \left[m_2^i m_2^i \right] - \left[m_1^i m_2^i \right]^2, \qquad (11)$$

$$\Delta_{1} = \begin{bmatrix} y_{i}m_{1}^{i} \end{bmatrix} \begin{bmatrix} m_{2}^{i}m_{2}^{i} \end{bmatrix} - \begin{bmatrix} y_{i}m_{2}^{i} \end{bmatrix} \begin{bmatrix} m_{1}^{i}m_{2}^{i} \end{bmatrix}, \qquad (12)$$

$$\Delta_2 = \begin{bmatrix} y_i m_2^i \end{bmatrix} \begin{bmatrix} m_1^i m_1^i \end{bmatrix} - \begin{bmatrix} y_i m_1^i \end{bmatrix} \begin{bmatrix} m_1^i m_2^i \end{bmatrix}.$$
(13)

Collection Czechoslov, Chem, Commun. /Vol. 37/ (1972)

By arrangement of these expressions we get $\Delta = x^2 A_{1+}x^3 A_{2+} \dots$, $\Delta_1 = x B_1 + x^2 B_{2+} \dots$, $\Delta_2 = x C_1 + x^2 C_{2+} \dots$, where A_j , B_j , C_j , j = 1,2 are independent of x, and $C_1 = -B_1$. As a consequence, for $x \to 0$ both coefficients M_1 , M_2 approach infinity with opposite signs. In contrast, the limit for the sum of squares is finite.

Cases II and III give both the same results: The value of M_1 goes to infinity for vanishing x. However, the formulas for S give a finite value. From the above results it further follows that in the neighbourhood of $m_2 \rightarrow m_1$ the approximation formulas (I), (4) and (7) in the given form lose their significance. Also the numerical computations for m_2 close to m_1 carry great errors since the expression $M_1m_1^i + M_2m_2^i$ represents the difference of two almost same quantities because $M_1 \approx -M_2$. Thus in the computation as many as 4, 5 or even more orders of accuracy are lost. A substitute for these formulas, which does not lose its meaning for $m_2 \rightarrow m_1$ and gives accurate results, is found similarly: Substituting $m_2 = m_1(1 + x) = m(1 + x)$ and in all expansions retaining only the terms with x. For the general Case I the substitutive formula is

$$u_{i} = (N_{1} + iN_{2}) m^{i}, \qquad (14)$$

or generally for $x = x_0 + ih$

$$u = (R_1 + xR_2) p^x . (15)$$

In special cases $y_0 = 0$ and $y_0 = 1$ we get for i = 0

$$u_i = iN_2m^i$$
 and $u_i = (1 + iN_2)m^i$. (16, 17)

The other extreme, when m_2 vanishes poses no numerical difficulties in contrast to the previous one. It should be noted that the approximation formulas (1), (4) and (7) do not transform into a single exponential, since $\lim m_2^i = 0$, for $m_2 = 0$ and i > 0, but $\lim m_2^i = 1$, for i = 0. Thus in the limit $m_2 = 0$ the formulas (1), (4)and (7) take the following form

$$u_0 = M_1 + M_2, \quad u_i = M_1 m_1^i \quad \text{for} \quad i > 0,$$
 (1a)

$$u_0 = 0, \quad u_i = M_1 m_1^i \quad \text{for} \quad i > 0,$$
 (4a)

$$u_0 = 1$$
, $u_i = M_1 m_1^i$ for $i > 0$. (7a)

One of the supposition here and in the earlier paper¹ was the assumption of an equal spacing of R pairs of experimental data (y_i, x_i) , $i = 0, 1 \dots R - 1$, *i.e.* $x_i = x_0 + ih$, where h is the increment of the argument. Let us assume now that the distribution of the values of x_i along the x axis is arbitrary. We shall perform the calculation

under these more general conditions for one special case of the approximation formula of the following type

$$u = P_1 \exp(\omega_1 x) + P_2 \exp(\omega_2 x), \qquad (18)$$

which can be modified by introducing $m_i = \exp(\omega_i)$, j = 1, 2, to give

$$u = P_1 m_1^{\mathbf{x}} + P_2 m_2^{\mathbf{x}} \,. \tag{19}$$

Let m_1 and m_2 be known (e.g. by an arbitrary estimate). The coefficients P_1 and P_2 are calculated from the equations

$$P_1[11] + P_2[12] = [y1], \quad P_1[12] + P_2[22] = [y2]$$
(20)

following from the conditions

$$\partial S / \partial \omega_1 = \partial S / \partial \omega_2 = 0 \tag{21}$$

of the minimum sum of square deviations

$$S = \sum_{i=0}^{R-1} (u_i - y_i)^2 .$$
 (22)

For brevity we introduced in Eqs (20) the following notation

$$[jk] = \sum_{i=0}^{R-1} m_j^{x_i} m_k^{x_i}, \quad [yj] = \sum_{i=0}^{R-1} y_i m_j^{x_i}, \quad (23)$$

where j, k = 1, 2. Substituting now m_1 and m_2 by $m_1(1 + r_1)$ and $m_2(1 + r_2)$, and taking m_j as constants and r_j as variables, we can express both coefficients $P_j, j = 1, 2$, and the sum of square deviations (Eq. (22)) as functions of r_1 and r_2 . Providing that r_1 and r_2 are sufficiently small we can neglect higher powers of r_j in the expansion of $(1 + r_j)$ and write

$$(1+r_j)^{x_i} = 1 + {x_i \choose 1} r_j + {x_i \choose 2} r_j^2 + {x_i \choose 3} r_j^3, \qquad (24)$$

for $j = 1, 2; i = 0, 1, \dots R - 1$.

Introducing the polynomials

$$\begin{split} L_0 &= \mathcal{Q}^{\mathbf{x}_0} + \mathcal{Q}^{\mathbf{x}_1} + \ldots + \mathcal{Q}^{\mathbf{x}_N} ,\\ L_1 &= \begin{pmatrix} x_0 \\ 1 \end{pmatrix} \mathcal{Q}^{\mathbf{x}_0} + \begin{pmatrix} x_1 \\ 1 \end{pmatrix} \mathcal{Q}^{\mathbf{x}_1} + \ldots + \begin{pmatrix} x_N \\ 1 \end{pmatrix} \mathcal{Q}^{\mathbf{x}_N} ,\end{split}$$

Collection Czechoslov. Chem. Commun. /Vol. 37/ (1972)

$$L_{2} = {\binom{x_{0}}{1}}^{2} \mathcal{Q}^{x_{0}} + {\binom{x_{1}}{1}}^{2} \mathcal{Q}^{x_{1}} + \dots + {\binom{x_{N}}{1}}^{2} \mathcal{Q}^{x_{N}},$$

$$L_{3} = {\binom{x_{0}}{2}} \mathcal{Q}^{x_{0}} + {\binom{x_{1}}{2}} \mathcal{Q}^{x_{1}} + \dots + {\binom{x_{N}}{2}} \mathcal{Q}^{x_{N}},$$

$$L_{4} = {\binom{x_{0}}{1}} {\binom{x_{0}}{2}} \mathcal{Q}^{x_{0}} + {\binom{x_{1}}{1}} {\binom{x_{1}}{2}} \mathcal{Q}^{x_{1}} + \dots + {\binom{x_{N}}{1}} {\binom{x_{N}}{2}} \mathcal{Q}^{x_{N}},$$

$$L_{5} = {\binom{x_{0}}{3}} \mathcal{Q}^{x_{0}} + {\binom{x_{1}}{3}} \mathcal{Q}^{x_{1}} + \dots + {\binom{x_{N}}{3}} \mathcal{Q}^{x_{N}}, \quad N = \mathbb{R} - 1,$$
(25)

in which $Q = m_j m_k$, j, k = 1, 2, we can easily express [jk] as a polynomial of the 3rd degree in r_1 and r_3 . Thus approximately

$$[jk] = L_0 + L_1(r_j + r_k) + L_2r_jr_k + L_3(r_j^2 + r_k^2) + L_4(r_j^2r_k + r_jr_k^2) + L_5(r_j^3 + r_k^3).$$
(26)

Similarly for j = 1, 2

$$\begin{bmatrix} yj \end{bmatrix} = \sum_{i=0}^{N} y_i m_j^{x_i} (1+r_j)_j^{x_i} = y_0 m_j^{x_0} + y_1 m_j^{x_1} + \dots + y_n m_j^{x_N} +$$
(27)
+ $r_j \begin{bmatrix} \binom{x_0}{1} y_0 m_j^{x_0} + \binom{x_1}{1} y_1 m_j^{x_1} + \dots + \binom{x_N}{1} y_N m_j^{x_N} \end{bmatrix} +$
+ $r_j^2 \begin{bmatrix} \binom{x_0}{2} y_0 m_j^{x_0} + \dots + \binom{x_N}{2} y_N m_j^{x_N} \end{bmatrix} +$
+ $r_j^3 \begin{bmatrix} \binom{x_0}{3} y_0 m_j^{x_0} + \dots + \binom{x_N}{3} y_N m_j^{x_N} \end{bmatrix} +$

Thus the problem formally becomes that treated earlier since all sums [jk] and [yj] were expressed as polynomials in r_1 and r_2 . The sum of square deviations can also be found by the earlier described method as a polynomial having for instance the following form

$$S = S_0 + S_1 r_1 + S_2 r_2 + \dots + S_8 r_1^3 + S_9 r_2^3.$$
⁽²⁸⁾

To this polynomial we then apply the iteration procedure described earlier¹ for finding the corrections r_1 and r_2 .

REFERENCES

Slavíček E.: This Journal 35, 2885 (1970).

Translated by V. Staněk.

Collection Czechoslov. Chem. Commun. /Vol. 37/ (1972)